Asian American Access to Care in the Affordable Care Act Era: Findings from a Population-Based Survey in California

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BACKGROUND: Though Asian Americans made gains in coverage following the Affordable Care Act (ACA), substantial variations in access to care remain across different ethnic subgroups. Several states are considering adoption of policies to collect health data for Asian Americans that is disaggregated by ethnic subgroup, which may identify disparities in access to care.

OBJECTIVE: We examined coverage and access to care between non-Hispanic White and Asian American adults following the ACA in California. We first compared outcomes in non-Hispanic White adults with all Asian Americans in our sample, and then evaluated whether we detect disparities when data is disaggregated into five of the most populous ethnic subgroups (Chinese, Korean, Filipino, Vietnamese, and Japanese).

DESIGN: Cross-sectional California Health Interview Survey data were collected between January 2014 and December 2016.

PARTICIPANTS: Our sample included 19,201 non-Hispanic White and 3077 Asian American non-elderly adults age 18 to 64 living in California.

MAIN MEASURES: Our outcomes were (1) being uninsured, (2) having a usual source of care, (3) delaying necessary medical care, and (4) delaying necessary prescription medications. Using multivariable logistic regression models, we examined our outcomes, adjusting for predisposing, enabling, need, and acculturation factors.

KEY RESULTS: Compared with non-Hispanic Whites, some subgroups of Asian Americans reported significantly worse access to care: disaggregated, adjusted analyses revealed that Koreans were significantly less likely to report a usual source of care (adjusted odds ratio [AOR] = 0.31, p < 0.01) relative to non-Hispanic Whites. Chinese (AOR = 0.42, p < 0.01) and Vietnamese (AOR = 0.34, p < 0.01) adults were significantly less likely to delay necessary care. **CONCLUSIONS:** Disaggregated analyses identified differences in access to care for Asian American subgroups following the ACA. State policies to collect disaggregated health data for Asian Americans may reveal heterogeneity in experiences of care and inform specific policies to reduce disparities in access to care.

KEY WORDS: access to care; disparities; minority health; health care reform; health policy.

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INTRODUCTION

Approximately 17.8 million individuals in the USA, comprising 6% of the population, self-identify as Asian American.¹ One of the fastest growing populations in the country, Asian Americans are a diverse race, encompassing more than 50 ethnicities and 100 languages spoken.^{2, 3} The sociodemographic heterogeneity of this racial group has led to efforts that standardize disaggregation of Asian American health data in state and federal surveys over the last decade; however, many of these items have not been widely adopted and variations in access to care among Asian American subgroups remain unknown.^{4, 5} The Office of Minority Health developed standards for disaggregating Asian American data as part of the Affordable Care Act (ACA), resulting in more granular data collection for some groups (e.g., Asian Indians, Chinese, Filipinos, Japanese, Koreans, Vietnamese), but not other groups (e.g., Bangladeshi, Burmese).²

Evidence suggests that Asian Americans have worse selfreported mental health⁶ and higher rates of chronic conditions,^{7, 8} compared with non-Hispanic Whites. However, despite these documented disparities and unique challenges in navigating the health care system-including language barriers and health literacy-there is limited recent evidence regarding access to care among Asian Americans.^{9, 10} Many prior studies have aggregated outcomes for the entire Asian American population; aggregated analyses may fail to detect important disparities for some subgroups.^{11–14} Further, they have used data predating the implementation of the ACA, which attempted to reduce disparities in access to and affordability of care through expansion of Medicaid eligibility, availability of parental coverage for dependents 26 and under, and reducing cost-sharing for routine preventive visits. Asian Americans made sizable gains in coverage following the insurance expansions in 2014, but substantial variations in rates in uninsurance remained across subgroups.1, 15-20

We examined differences in coverage and access to care between non-Hispanic White and Asian American non-elderly adults in California 3 years after the ACA, between 2014 and 2016. To assess what differences may or may not be captured by aggregated data, we first compared outcomes in non-



Hispanic White adults with all Asian Americans (aggregated) and then examined outcomes for five of the most populous subgroups in California: Chinese, Korean, Filipino, Vietnamese, and Japanese.

METHODS

Data

The nation's largest state health survey, California Health Interview Survey (CHIS) is a state-representative, randomdigit-dial telephone survey administered annually to about 20,000 households by the University of California, Los Angeles Center for Health Policy Research in collaboration with the California Department of Public Health and Department of Health Care services.²¹ Its sampling frame uses both landline and cell phone random-digit-dialing techniques, as well as a supplemental surname list to ensure accurate representation from Korean and Vietnamese households.

CHIS is also one of the few publicly available, federally or state-administered surveys that asks about specific ethnic subgroup among Asian American respondents. To accurately reflect the diversity of the state, CHIS is conducted in multiple languages including Chinese (Cantonese and Mandarin dialects), Korean, Tagalog, and Vietnamese.

Outcomes

Our outcomes were (1) being uninsured, (2) having a usual source of care, (3) delaying necessary medical care in the past 12 months, and (4) delaying or not getting medication prescribed by a doctor in the last 12 months. As secondary outcome, we also examined whether respondents reported a clinic or health center as their usual site of care, among those who reported having a usual source of care.

Study Sample

We pooled 3 years of CHIS data collected between January 2014 and December 2016. Our sample included adults age 18-64 who self-identified as either non-Hispanic White or non-Hispanic Asian American residing in California. Given our focus on insurance coverage and access to care following the ACA, we excluded adults age 65 or older, because Medicare provides nearly universal coverage for this group. Consistent with previous studies, we excluded respondents reporting other races and ethnicities (Hispanic or Latino, Black or African American, individuals reporting multiple races, and Asian respondents who reported more than one Asian ethnicity).¹⁵ We also excluded respondents who self-identified as "Other Asian," a category that is comprised of Cambodian, Hmong, Indian, Indonesian, Laotian, Malaysian, Pakistani, Sri Lankan, Taiwanese, Thai, and other subgroups. While it is unclear how these groups are distributed within this category, this

aggregation likely indicates a small number of respondents in each subgroup, and an opportunity for future analyses was further disaggregated data to become available. We grouped respondents into six mutually exclusive categories: White, Chinese, Korean, Filipino, Vietnamese, and Japanese.

Andersen Behavioral Model for Health Services Use

The Andersen behavioral model is a conceptual framework that conceives access to and utilization of care as a function of three main factors: an individual's beliefs or propensity (predisposing factors), resources (enabling factors), and perceived need for health services.²² Iterations of the Andersen model have integrated social or structural determinants of access that are more racially or ethnically specific, with some studies identifying acculturation factors (e.g., English proficiency, length of residence) as key drivers of differences in access to care between Asian Americans and Whites.^{18–20, 23–25}

Statistical Analysis

We used Pearson's chi-square tests to describe differences first between White and all Asian Americans in our sample (aggregated), and then compared differences across White and the five subgroups (disaggregated). We also examined the characteristics of respondents who identified as "Other Asian" or who reported more than one Asian subgroup, which we present in a supplementary analysis. We first calculated unadjusted rates of each of our outcomes. We then estimated logistic regression models across our outcomes using the Andersen behavioral model: we adjusted for predisposing factors (e.g., age, gender, highest level of education attained, marital status, and household size), enabling factors (e.g., household income measured as a percentage of the federal poverty level, employment status, urban/rural designation, and being uninsured), need factors (e.g., self-reported health status), and acculturation factors (e.g., being born in the USA, percent of life spent in the USA, citizenship status, and English proficiency).²⁶⁻²⁸ To examine the degree to which these characteristics explained differences in access to care between White and Asian American respondents, we constructed five models: model 1 was unadjusted, model 2 included predisposing characteristics, model 3 added enabling factors, model 4 added need factors, and model 5 added acculturation factors. As a sensitivity analysis, we also included respondents who identified as "Other Asian" or reported more than one subgroup in our aggregated group. Consistent with previous studies examining disaggregated Asian subgroups, we applied a more conservative alpha level (p < 0.01) to account for multiple comparisons.29, 30

Our regression models used jackknife replication variance estimates and applied survey weights to account for the complex sampling design of CHIS. All analyses were conducted in Stata 15 (StataCorp LP, College Station, TX, USA).

RESULTS

Our unweighted sample consisted of 22,098 respondents, of which 19,021 were White and 3077 were Asian American: 1247 were Chinese, 327 were Korean, 560 were Filipino, 585 were Vietnamese, and 358 were Japanese. Our weighted sample characteristics comparing White respondents and all Asian American respondents (aggregate) and disaggregated into five subgroups are presented in Table 1. Relative to White respondents, significantly more Asian American respondents reported having a college degree or higher (59.00% vs. 48.90%, p < 0.01); however, significantly fewer Asian American respondents reported earning more than 400% of FPL (48.23% vs. 56.05%, p < 0.01).

When disaggregated into five subgroups, significantly fewer Vietnamese respondents reported having a college degree or higher (39.00%, p < 0.01) relative to White respondents (48.90%). There was wide variation in income: more than one-

Table 1 Characteristics of Non-Hispani	: White and Asian Respondents.	Disaggregated by	Subgroup, 2014–2016

	White <i>n</i> = 19,021	Asian Ameri (aggre n=307	gate)	Chines $n = 124$		Korean n = 327		Filipin <i>n</i> = 560		Vietna <i>n</i> = 585		Japano n = 358	
		%	<i>p</i> *	%	<i>p</i> *	%	<i>p</i> *	%	<i>p</i> *	%	<i>p</i> *	%	<i>p</i> *
Predisposing factors													
Age 18–34	32.44	40.89	<	44.22	<	46.19	0.03	40.82	0.05	35.13	0.56	25.53	0.32
35-64	52.44 67.66	40.89 59.11	0.01	44.22 55.78	0.01	53.81	0.05	40.82 59.18	0.05	64.87	0.50	23.33 74.47	0.52
Female	49.21	54.22	<	53.43	0.13	61.39	0.03	52.32	0.37	51.53	0.63	61.45	0.10
	.,	0 1122	0.01	00110	0110	01107	0.00	02102	0107	01100	0102	01110	0110
Education	2.05	6 70		7.00		2 00	0.16	0.07	0.05	00.10		0.00	0.01
Does not have high school diploma	3.95	6.78	< 0.01	7.22	< 0.01	2.99	0.16	0.86	0.05	23.13	< 0.01	0.88	0.01
High school diploma	19.33	15.38	0.01	12.09	0.01	16.82		15.20		23.38	0.01	10.66	
Some college	15.74	12.11		7.98		11.06		18.82		10.27		5.95	
Vocational school/AA or	12.09	6.73		4.23		6.54		9.49		4.21		13.53	
AS degree	12.09	0.75		1.20		0.01		2.12		1.21		10.00	
College degree or higher	48.90	59.00		68.47		62.60		55.64		39.00		68.98	
Marital status: married	51.68	53.38	0.42	51.87	0.96	54.24	0.67	50.49	0.78	59.18	0.11	61.02	0.14
Enabling factors													
Income, by federal poverty l													
0–138% FPL	14.34	20.24	<	19.30	0.23	16.61	0.31	16.94	0.07	36.93	<	6.29	0.35
139–249% FPL	12.27	13.86	0.01	10.89		17.18		16.68		13.25	0.01	9.97	
250–399% FPL	17.34	17.66		16.39		20.65		19.84		13.63		17.43	
400% FPL or more	56.05	48.23	0.40	53.41	0.00	45.56	0.77	46.54	0.05	36.19	0.50	66.32	0.00
Unemployed	25.48	23.85	0.40	22.87	0.39	27.13	0.77	22.20	0.35	27.89	0.56	20.53	0.33
Health insurance coverage Uninsured	7.28	10.26	0.01	9.03	0.32	10.48	0.20	10.62	0.12	12.86	0.02	8.45	0.75
Employer-sponsored	62.84	58.63	0.01	61.33	0.32	47.63	< 0.20	63.93	0.12	46.93	< 0.02	69.30	0.73
Employer-sponsored	02.84	58.05	0.05	01.55	0.05	47.05	0.01	05.95	0.78	40.95	0.01	09.50	0.27
Medicaid	14.61	18.79	0.01	14.34	0.91	19.09	0.31	19.84	0.08	29.47	<	9.04	0.30
	1 1101	101/2	0.01	1 110 1	0191	19109	0.01	19101	0.00		0.01	2.0.	0.00
Individually purchased	10.92	10.55	0.79	14.11	0.17	20.41	<	2.99	<	9.81	0.67	12.48	0.67
v 1							0.01		0.01				
Medicare	4.36	1.76	<	1.18	<	2.39	0.42	2.62	0.12	1.12	<	0.73	<
			0.01		0.01						0.01		0.01
Need factors		1= 0.6				1							0.00
Fair or poor self-reported	13.14	17.86	<	15.82	0.28	17.38	0.31	14.46	0.59	30.61	<	15.14	0.69
health	(5.72	40.14	0.01	42.07		21.00		66.12	0.02	2776	0.01	52.05	0.10
One or more chronic conditions	65.73	49.14	< 0.01	43.97	< 0.01	31.98	< 0.01	66.13	0.03	37.76	< 0.01	52.85	0.10
Acculturation factors			0.01		0.01		0.01				0.01		
Born in the USA	89.78	32.16	<	30.01	<	24.18	<	36.49	<	20.81	<	70.02	<
Doin in the 0.57	07.70	52.10	0.01	50.01	0.01	21.10	0.01	50.17	0.01	20.01	0.01	70.02	0.01
US citizen	96.55	77.73	<	75.01	<	60.17	<	81.66	<	87.81	<	82.72	<
			0.01		0.01		0.01		0.01		0.01		0.01
Percent of life in the USA													
0–20	1.68	14.17	<	18.92	<	14.78	<	11.29	<	12.57	<	5.03	<
21-40	1.89	14.89	0.01	17.08	0.01	16.78	0.01	13.17	0.01	15.20	0.01	5.98	0.01
41-60	2.25	17.78		16.31		16.89		17.52		25.08		9.68	
61-80	1.95	12.18		12.21		13.21		8.56		21.16		5.34	
81 or more	92.22	41.01		35.48		38.33		49.46		25.99		73.96	
English proficiency	99.68	82.22	< 0.01	78.89	< 0.01	74.13	< 0.01	97.26	< 0.01	60.51	< 0.01	94.81	< 0.01
			0.01		0.01		0.01		0.01		0.01		0.01

*p value according to χ^2 tests for differences, relative to non-Hispanic White respondents. Survey weights were applied to reflect population distributions

third of Vietnamese respondents reported income of 0–138% of the federal poverty level, more than double the proportion of White respondents (14.34%, p < 0.01). Significantly more Vietnamese respondents (29.47%) reported Medicaid coverage relative to their White peers (14.61%, p < 0.01), while more Korean (20.41%) respondents reported individually purchased insurance relative to Whites (10.92%, p < 0.01). In Appendix Table 6, we report characteristics of respondents who reported "Other Asian" ethnicity or who reported more than one Asian ethnicity (e.g., Chinese and Vietnamese).

Table 2 summarizes the association between race or ethnic subgroup and being uninsured. Both in the aggregate and disaggregated, there were no statistically significant differences in uninsurance between Asian Americans and White respondents.

As shown in Table 3, significantly fewer Asian Americans had a usual source of care relative to their White peers in the aggregate (80.02% vs. 88.06%, p < 0.01), as were Korean (63.85%, p < 0.01) and Vietnamese (77.97%, p < 0.01) respondents when disaggregated. In the aggregate, this disparity persisted across our models that adjusted for predisposing, enabling, and need factors, but was no longer significant when acculturation factors were included (model 5, AOR = 0.76, p = 0.11). Vietnamese respondents were no longer significantly less likely to have a usual source of care after adjusting for predisposing (model 2, AOR = 0.47, p = 0.02); however, Koreans were significantly less likely to have a usual source of care across all adjusted models (model 5, AOR = 0.31, p < 0.01).

In Table 4, significantly fewer Asian Americans (aggregate) reported delaying necessary medical care compared with White respondents (9.31% vs. 18.32%, p < 0.01) in unadjusted analyses—in particular, Chinese (7.88%, p < 0.01), Filipino (8.83%, p < 0.01), and Vietnamese (6.72%, p < 0.01). In terms of magnitude and significance, these associations in both the aggregate and for Chinese and Vietnamese respondents persisted as models included predisposing, enabling, need, and acculturation factors. The association was no longer significant for Filipino respondents upon inclusion of acculturation factors (AOR = 0.43, p = 0.01). Similarly, Asian

Americans (aggregate) were significantly less likely to report delaying or not getting prescription medications relative to White respondents (6.31% vs. 12.71%, p < 0.01) in unadjusted analyses, as presented in Table 5. In disaggregated, unadjusted analyses, Chinese (3.91%, p < 0.01) respondents were significantly less likely to report delaying or not getting prescription medications. Estimates were comparable in magnitude and significance when incrementally adjusting for predisposing, enabling, need, and acculturation factors.

Among those who reported a usual source of care, Asian Americans (aggregate) were significantly more likely to be going to a clinic or health center relative to White respondents (32.03% vs. 21.23%, p < 0.01). Chinese (38.20%, p < 0.01) and Filipino (37.01%, p < 0.01) respondents were more likely to report a clinic or health center as their usual source of care (Appendix Table 7) compared with White respondents. In terms of magnitude and significance, these associations remained across our five models.

Sensitivity analyses that included those who identified as "Other Asian" or who reported multiple subgroups did not alter our aggregated results in terms of significance or magnitude for our outcomes.

DISCUSSION

We identified significant differences in access to care between White and Asian American non-elderly adults in California, a finding that persisted (1) in unadjusted and adjusted analyses and (2) in an aggregate Asian American group and when disaggregated by subgroup. Direction and magnitude of associations, however, varied, reflecting heterogeneity in the experiences of access to care across subgroups. Previous studies suggest significant associations between access to care and measures of acculturation, such as limited English proficiency, citizenship, and time spent in the USA and access to care.^{28, 31, 32} Our study builds upon this work by using contemporary data after ACA coverage expansions for five Asian subgroups. For some of our

Race/ethnic subgroup	Unadjusted rates		Model 1: unadjusted		Model 2*: predisposing		Model 3 [†] : predisposing, enabling		Model 4 [‡] : predisposing, enabling, need		Model 5 [§] : predisposing, enabling, need, acculturation	
	%	р	OR	р	AOR	р	AOR	р	AOR	р	AOR	р
Non-Hispanic White (ref)	7.28	_	1.00	_	1.00	-	1.00	-	1.00	-	1.00	_
Asian Americans (aggregate)	10.26	0.01	1.46	0.01	1.62	< 0.01	1.45	0.03	1.48	0.02	1.12	0.58
Chinese	9.03	0.32	1.26	0.33	1.41	0.17	1.28	0.32	1.28	0.31	0.86	0.63
Korean	10.48	0.20	1.49	0.21	1.75	0.11	1.49	0.26	1.47	0.28	0.93	0.84
Filipino	10.62	0.12	1.51	0.12	1.62	0.09	1.48	0.17	1.51	0.15	1.25	0.46
Vietnamese	12.68	0.02	1.85	0.02	2.01	0.02	1.69	0.06	1.79	0.04	1.37	0.35
Japanese	8.45	0.75	1.18	0.74	1.50	0.40	1.55	0.36	1.62	0.31	1.45	0.45

Table 2 Association Between Race and Uninsurance Among Non-Hispanic White and Asian Adults Age 18-64 in California, 2014-2016

*Predisposing factors were age, gender, highest level of education attained, marital status, and household size

[†]Enabling factors were household income measured as a percentage of the federal poverty level, employment status, and urban/rural designation ${}^{\sharp}$ Need factor was self-reported health status

 s Acculturation factors were being born in the USA, being a US citizen, percent of life spent in the USA, and English proficiency

Race/ethnic subgroup	Unadjusted rates		Model 1: unadjusted		Model 2*: predisposing		Model 3 [†] : predisposing, enabling		Model 4 [‡] : predisposing, enabling, need	Model 5 [§] : predisposing, enabling, need, acculturation		
	%	р	OR	р	AOR	р	AOR	р	AOR	р	AOR	р
Non-Hispanic White (ref)	88.06	_	1.00	_	1.00	_	1.00	_	1.00	_	1.00	_
Asian Americans (aggregate)	80.02	< 0.01	0.54	< 0.01	0.55	< 0.01	0.61	< 0.01	0.60	< 0.01	0.76	0.11
Chinese	81.84	0.01	0.61	0.02	0.66	0.05	0.72	0.14	0.70	0.12	0.91	0.73
Korean	63.85	< 0.01	0.24	< 0.01	0.22	< 0.01	0.23	< 0.01	0.23	< 0.01	0.31	< 0.01
Filipino	85.24	0.24	0.78	0.24	0.79	0.28	0.89	0.57	0.88	0.53	1.07	0.77
Vietnamese	77.97	< 0.01	0.48	< 0.01	0.47	0.02	0.58	0.09	0.54	0.06	0.65	0.22
Japanese	80.87	0.12	0.57	0.12	0.47	0.05	0.48	0.04	0.46	0.04	0.50	0.06

 Table 3 Association Between Race and Reporting a Usual Source of Care Among Non-Hispanic White and Asian Adults Age 18–64 in California, 2014–2016

* Predisposing factors were age, gender, highest level of education attained, marital status, and household size

[†]Enabling factors were household income measured as a percentage of the federal poverty level, employment status, urban/rural designation, and being uninsured [†]Need factor was self-reported health status

 $^{\$}$ Acculturation factors were being born in the USA, being a US citizen, percent of life spent in the USA, and English proficiency

outcomes, disparities between Whites and Asian Americans—both aggregated and disaggregated—attenuated upon inclusion of acculturation factors. These findings, supported by previous studies, suggest that acculturation mediates the relationship between ethnicity and access to care.^{18, 20, 33} The impact of acculturation may vary by Asian subgroup, further supporting the importance of disaggregated analyses.

Uninsured

Before the ACA, Asian Americans reported higher uninsurance rates relative to White adults, a gap attributed in part to higher rates of employment in small businesses that did not offer health insurance benefits.³⁴ Studies predating the ACA identified variation in the sociodemographic and acculturationrelated characteristics associated with uninsurance across Asian American subgroups in California.¹⁷ Following the ACA, disparities in insurance coverage between White and Asian American respondents appear to have attenuated, perhaps reflecting targeted outreach—which relied on a network of hundreds of community organizations, churches, and small businesses—and availability of in-language assistance for enrollment in California's marketplace, Covered California.^{15, 35, 36} In the first few months of the ACA implementation, more than 20% of new Covered California enrollees were of Asian descent. An estimated 57% of Chinese, 65% of Vietnamese, and 70% of Korean new enrollees in Covered California plans did so through insurance agents.³⁷

Our results also suggest that, following the ACA, there was heterogeneity in the distribution of insurance types across Asian American subgroups, a finding consistent with a study using 2003 and 2005 CHIS data.¹⁷ In comparison with previous estimates using CHIS, we found that there were decreases in the proportion of Asian Americans (aggregated and disaggregated) who were uninsured, and there were increases in having individually purchased and public insurance in some subgroups following ACA implementation.

Usual Source of Care

In studies predating the ACA, Asian Americans in California were significantly less likely to have a usual source of care

 Table 4 Association Between Race and Delaying Medical Care in the Past 12 Months Among Non-Hispanic White and Asian Adults Age 18–64 in California, 2014–2016

Race/ethnic subgroup	Unadjusted rates		Model 1: unadjusted		Model 2*: predisposing		Model 3 [†] : predisposing, enabling		Model 4 [‡] : predisposing, enabling, need		Model 5 [§] : predisposing, enabling, need, acculturation	
	%	р	OR	р	AOR	р	AOR	р	AOR	р	AOR	р
Non-Hispanic White (ref) Asian Americans (aggregate) Chinese Korean Filipino Vietnamese Japanese	18.32 9.31 7.88 17.39 8.83 6.72 9.92	- <0.01 <0.01 0.86 <0.01 <0.01 0.09	$1.00 \\ 0.46 \\ 0.38 \\ 0.94 \\ 0.43 \\ 0.32 \\ 0.49$	- <0.01 <0.01 0.86 <0.01 <0.01 0.11	1.00 0.46 0.38 0.97 0.43 0.32 0.47	- <0.01 <0.01 0.93 <0.01 <0.01 0.10	$1.00 \\ 0.44 \\ 0.36 \\ 0.91 \\ 0.41 \\ 0.30 \\ 0.47$	- < 0.01 < 0.01 0.79 < 0.01 < 0.01 0.10	$ \begin{array}{c} 1.00\\ 0.40\\ 0.33\\ 0.86\\ 0.38\\ 0.25\\ 0.44 \end{array} $	- <0.01 <0.01 0.68 <0.01 <0.01 0.07	$ \begin{array}{c} 1.00\\ 0.48\\ 0.42\\ 1.05\\ 0.43\\ 0.34\\ 0.44 \end{array} $	- < 0.01 < 0.01 0.90 0.01 < 0.01 0.08

*Predisposing factors were age, gender, highest level of education attained, marital status, and household size

[†]Enabling factors were household income measured as a percentage of the federal poverty level, employment status, urban/rural designation, and being uninsured [‡]Need factor was self-reported health status

 s Acculturation factors were being born in the USA, being a US citizen, percent of life spent in the USA, and English proficiency

Table 5 Association Between Race and Delaying or Not Filling Prescription Medication in the Past 12 Months Among Non-Hispanic White and
Asian Adults Age 18–64 in California, 2014–2016

Race/ethnic subgroup	Unadjusted rates		Model 1: unadjusted		Model 2*: predisposing		Model 3 [†] : predisposing, enabling		Model 4 [‡] : predisposing, enabling, need		Model 5 [§] : predisposing, enabling, need, acculturation	
	%	р	OR	р	AOR	р	AOR	р	AOR	р	AOR	р
Non-Hispanic White (ref) Asian Americans (aggregate) Chinese Korean Filipino Vietnamese Japanese	12.71 6.31 3.91 2.94 10.16 6.16 6.64	- < 0.01 < 0.01 < 0.01 0.36 0.05 0.09	$ \begin{array}{r} 1.00\\ 0.46\\ 0.28\\ 0.21\\ 0.78\\ 0.45\\ 0.49 \end{array} $	- < 0.01 < 0.01 0.02 0.36 0.06 0.10	$ \begin{array}{r} 1.00\\ 0.48\\ 0.30\\ 0.22\\ 0.82\\ 0.44\\ 0.49 \end{array} $	- < 0.01 < 0.01 0.03 0.47 0.06 0.10	$ \begin{array}{c} 1.00\\ 0.47\\ 0.29\\ 0.21\\ 0.80\\ 0.41\\ 0.49 \end{array} $	- < 0.01 < 0.01 0.03 0.42 0.05 0.11	1.00 0.41 0.25 0.19 0.72 0.33 0.44	- < 0.01 < 0.01 0.02 0.25 0.02 0.08	1.00 0.52 0.30 0.23 0.86 0.39 0.48	- 0.01 < 0.01 0.03 0.63 0.09 0.12

*Predisposing factors were age, gender, highest level of education attained, marital status, and household size

[†]Enabling factors were household income measured as a percentage of the federal poverty level, employment status, urban/rural designation, and being uninsured [‡]Need factor was self-reported health status

[§]Acculturation factors were being born in the USA, being a US citizen, percent of life spent in the USA, and English proficiency

relative to Whites.^{19, 20} In the aggregate, our findings are consistent; however, we build upon these findings by disaggregating into five Asian American subgroups, wherein Korean respondents were significantly less likely to report having a usual source of care. A study using 2005 and 2009 CHIS data indicated Koreans were significantly less likely to report having a usual source of care relative to other Asian American subgroups, and this was attributed to higher rates of uninsurance. Our study suggests that, following the ACA, Koreans had similar rates of uninsurance relative to Whites, yet were still significantly less likely to report having a usual source of care. This finding persisted even after adjusting for predisposing, enabling, need, and acculturation factors. Vietnamese respondents were significantly less likely to report a usual source of care relative to White respondents; however, for Vietnamese respondents, the association was no longer significant after adjusting for predisposing factors.

Among those with a usual source of care, Asian Americans—specifically Chinese and Filipino respondents—were significantly more likely to go to a clinic or health center relative to Whites, which was consistent with previous studies.³⁸ Differences in site of usual source of care among Asian Americans have previously been explained in part by attitudes and perceptions about discrimination in health care.³⁸ That the site of care for some Asian American subgroups was a health center or clinic could reflect need for or availability of enabling services (e.g., on-site language interpretation services, transportation services, and culturally proficient care).^{11, 39}

Delays in Necessary Medical Care and Prescription Drugs

Our finding that Asian Americans in the aggregate were significantly less likely to report delays in necessary medical care or prescription drugs is consistent with a recent analysis, which suggested that, following ACA implementation, Asian Americans were significantly less likely to report delaying necessary care specifically because of cost.¹⁰ Our disaggregated

findings—wherein Chinese and Vietnamese respondents were significantly less likely to report delays in necessary medical care—align with a study preceding the ACA, and the authors suggested these findings reflect different cultural preferences and a more crisis-oriented perspective on access to care among Asian Americans.¹⁹ Our study builds upon this analysis by estimating the odds of delaying necessary care or prescription medication following the ACA, and by examining these outcomes for Japanese and Korean adults, who reported statistically comparable rates relative to their White peers.

State Policies in Disaggregation of Asian American Data

During the Obama administration, the White House Initiative on Asian Americans began to identify methods of collecting and reporting more detailed subgroup data in federal surveys.⁴⁰ There has been more movement at the state level: a bill requiring disaggregated data collection passed the California State Assembly and was signed into law in 2016.^{41, 42} Similar legislation passed in New York and Massachusetts the following year. Some civil rights groups and lawmakers argue that the disaggregated information may provide further information on disparities in poverty, education, and health care that currently go unnoticed in a diverse group and can inform community needs.^{43, 44} Opponents of the legislation suggest that the policy unfairly targets only Asian Americans, and that there are potential unintended consequences in disaggregation, particularly in terms of education.⁴⁵

Limitations

Our study did not estimate the causal impact of the Affordable Care Act on access to care for Asian Americans. The data in our study are limited to adults in California, and therefore our results may not be representative of outcomes in other states or nationwide. Notably, about 16% of the Californian population is Asian American, which is higher than the national average (6%).⁴⁶ Moreover, some counties in California expanded Medicaid eligibility before 2014, which we are unable to

account for in the data; nevertheless, the CHIS is one of the few publicly available data sources that has collected more granular Asian subgroup data for multiple years, and an estimated 31% of non-elderly Asian Americans in the USA reside in California.^{1, 47} Further, our sampling strategy excluded individuals who self-identified as Asian American but either identified as "Other Asian" or reported multiple subgroups, which prevented us from assessing further intra-racial heterogeneity. Examining access to care for multiracial populations is important for future studies. We ran sensitivity analyses to include respondents who reported multiple Asian ethnicities or identified as "Other Asian," in our aggregate analyses, which did not change the significance or direction for most of our findings. Lastly, our models currently compare differences between non-Hispanic White and Asian American nonelderly adults, and it is possible that non-Hispanic White adults would not be the appropriate comparator group.⁴⁸

Our findings identified critical differences in access to care between White and Asian American adults in California following the ACA, but these differences varied substantially among Asian American subgroups. These results suggest a potential benefit to collecting disaggregating data—particularly as it pertains to health care data—for Asian Americans, and can potentially inform more targeted public policy and program interventions to mitigate gaps in access to care.

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APPENDIX

Table 6 Respondent Characteristics, Comparing Non-Hispanic White and Other Asian Respondents in California, 2014–2016

Characteristic	White <i>n</i> = 19,021	Other <i>n</i> = 868	
		%	<i>p</i> *
Predisposing factors			
Age			
18–34	32.44	53.75	<
35-64	67.66	46.25	0.01
Female	49.21	48.16	0.71
Education			
Does not have high school diploma	3.95	3.46	<
High school diploma	19.33	19.44	0.01
Some college	15.74	9.73	
Vocational school/AA or AS degree	12.09	6.57	
College degree or higher	48.90	60.80	
Marital status: married	51.68	55.12	0.24
Enabling factors			
Income (by federal poverty level [FPL	D		
0–138% FPL	14.34	19.40	0.16
139–249% FPL	12.27	12.47	0.10
250–399% FPL	17.34	15.37	
400% FPL or more	56.05	52.75	
			0.25
Unemployed	25.48	28.40	0.35
Health insurance coverage	7.00	<i>((</i> 7	0.00
Uninsured	7.28	6.67	0.69
Employer-sponsored	62.84	63.55	0.84
Medicaid	14.61	19.58	0.04
Individually purchased	10.92	7.42	0.05
Medicare	4.36	2.78	0.19
Need factors			
Fair or poor self-reported health	13.14	10.42	0.25
One or more chronic conditions	65.73	57.63	<
			0.01
Acculturation factors			
Born in the USA	89.78	32.97	<
			0.01
US citizen	96.55	73.59	<
			0.01
Percent of life in the USA			
0-20	1.68	12.70	<
21-40	1.89	18.60	0.01
41-60	2.25	14.26	0.01
61-80	1.95	11.01	
81 or more	92.22	43.43	
English proficiency	92.22 99.68	94.59	<
English proficiency	J9.00	27.39	0.01
			0.01

*p value according to χ^2 tests for differences. Estimates were weighted to reflect population averages

Table 7 Association Between Race and Reporting a Clinic or Community Health Center as Usual Source of Care Among Non-Hispanic White and Asian Adults Age 18–64 in California, 2014–2016

Race/ethnic subgroup	Unadjusted rates				Model predisp	l 2*: Model sposing predisj enablin		osing,	Model 4 [‡] : predisposing, enabling, need		Model 5 [§] : predisposing, enabling, need, acculturation	
	%	р	OR	р	AOR	р	AOR	р	AOR	р	AOR	р
Non-Hispanic White (ref) Asian Americans (aggregate) Chinese Korean Filipino Vietnamese Japanese	21.23 32.03 38.20 22.19 37.01 19.31 15.87	- < 0.01 < 0.01 0.86 < 0.01 0.67 0.30	$ \begin{array}{r} 1.00\\ 1.75\\ 2.29\\ 1.06\\ 2.19\\ 0.89\\ 0.70\end{array} $	- <0.01 <0.01 0.86 <0.01 0.67 0.31	1.00 1.88 2.58 1.20 2.34 0.83 0.80	- < 0.01 < 0.01 0.56 < 0.01 0.53 0.55	1.00 1.92 2.72 1.23 2.35 0.72 0.96	- <0.01 <0.01 0.53 <0.01 0.29 0.92	1.00 1.90 2.69 1.22 2.32 0.71 0.95	- < 0.01 < 0.01 0.54 < 0.01 0.27 0.90	1.00 1.69 2.29 1.00 2.03 0.58 0.89	- < 0.01 < 0.01 0.99 < 0.01 0.17 0.77

*Predisposing factors were age, gender, highest level of education attained, marital status, and household size

Enabling factors were household income measured as a percentage of the federal poverty level, employment status, urban/rural designation, and being uninsured [‡]Need factor was self-reported health status

[§]Acculturation factors were being born in the USA, being a US citizen, percent of life spent in the USA, and English proficiency

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