

A NOTE CONCERNING BEHAVIOUR OF ITERATED LOGARITHM TYPE

C. C. HEYDE

Let $\{X_i, i=1, 2, 3, \dots\}$ be a sequence of independent and identically distributed random variables and write $S_n = \sum_{i=1}^n X_i, n \geq 1$. It is well known that

$$\Pr(\limsup_{n \rightarrow \infty} (2n \log \log n)^{-1/2} S_n = 1) = 1$$

if and only if $EX_i = 0, EX_i^2 = 1$ (Hartman and Wintner [6] obtained the sufficiency part and Strassen [9] the necessity) and the purpose of this note is to clarify the corresponding situation in the case where these moment conditions are violated. It turns out that the oscillation behaviour of normed sums rests essentially on whether or not the summands belong to the domain of partial attraction of the normal distribution.

Firstly, we remark that it has been shown by Freedman (see Strassen [9]) that there exist symmetric random variables X_i with $EX_i^2 = \infty$ or even $E|X_i| = \infty$, for which it is possible to choose a monotone sequence $\{b_n, n=1, 2, 3, \dots\}$ of positive constants with $b_n \rightarrow \infty$ as $n \rightarrow \infty$ such that

$$\Pr(\limsup_{n \rightarrow \infty} b_n^{-1} S_n = 1) = 1.$$

With this in mind, we go on to establish the following theorem.

THEOREM 1. *Suppose the X_i are such that there exists a monotone sequence of positive constants $\{b_n, n=1, 2, 3, \dots\}$ with $b_n \rightarrow \infty$ as $n \rightarrow \infty$ for which*

$$(1) \quad \limsup_{n \rightarrow \infty} b_n^{-1} |S_n| < \infty$$

with positive probability. Then, at least one of the following two conditions is satisfied:

(i) $b_n^{-1}(S_n - \text{med } S_n) \xrightarrow{\text{a.s.}} 0.$

(ii) *The X_i belong to the domain of partial attraction of the normal distribution. A necessary and sufficient condition for this is (Lévy [7, p. 113])*

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$$(2) \quad \liminf_{u \rightarrow \infty} \frac{u^2 \Pr(|X_1| > u)}{\int_{|x| \leq u} x^2 d \Pr(X_1 \leq x)} = 0.$$

("a.s." denotes almost sure convergence; "med" is an abbreviation of median).

PROOF. Firstly, since (1) is a tail event on the S_1, S_2, \dots process it has probability one using the Hewitt-Savage zero-one law (e.g. Breiman [1, Corollary 3.50]).

Next, suppose that neither (i) nor (ii) are satisfied. From (2), there exists a constant $C > 0$ such that

$$u^2 \Pr(|X_1| > u) > C \int_{|x| \leq u} x^2 d \Pr(X_1 \leq x)$$

for $u > U$. Thus, for $u > U$ and $\eta > 1$,

$$\begin{aligned} u^2 [\Pr(|X_1| > u) - \Pr(|X_1| > \eta u)] &\leq \int_{u < |x| \leq \eta u} x^2 d \Pr(X_1 \leq x) \\ &\leq u^2 C^{-1} \eta^2 \Pr(|X_1| > \eta u), \end{aligned}$$

from which we obtain

$$\Pr(|X_1| > u) \leq (1 + C^{-1} \eta^2) \Pr(|X_1| > \eta u).$$

Thus, for n sufficiently large we have, when $\epsilon > 1$,

$$(1 + C^{-1} \epsilon^2)^{-1} \Pr(|X_1| > b_n) \leq \Pr(|X_1| > \epsilon b_n) \leq \Pr(|X_1| > b_n),$$

and when $\epsilon < 1$,

$$\Pr(|X_1| > b_n) \leq \Pr(|X_1| > \epsilon b_n) \leq (1 + C^{-1} \epsilon^{-2}) \Pr(|X_1| > b_n),$$

so the convergence (divergence) of $\sum \Pr(|X_1| > \epsilon b_n)$ for some $\epsilon > 0$ implies convergence (divergence) for every $\epsilon > 0$.

If $\sum \Pr(|X_1| > \epsilon b_n) = \infty$ for every $\epsilon > 0$, we have from the Borel-Cantelli Lemmas that

$$(3) \quad \Pr(|X_n| > \epsilon b_n \text{ i.o.}) = 1$$

for every $\epsilon > 0$ ("i.o." is an abbreviation of infinitely often). Furthermore, $|S_n - S_{n-1}| = |X_n| \leq |S_n| + |S_{n-1}|$, so that (3) implies $\Pr(|S_n| > \frac{1}{2} \epsilon b_n \text{ i.o.}) = 1$ for every $\epsilon > 0$ which contradicts the condition (1).

Finally, suppose that $\sum \Pr(|X_1| > \epsilon b_n) < \infty$, every $\epsilon > 0$. We introduce a sequence of random variables $\{Y_k, k = 1, 2, 3, \dots\}$ defined by

$$Y_k = X_k \quad \text{if } |X_k| \leq b_k \\ = 0 \quad \text{if } |X_k| > b_k,$$

and write $T_n = \sum_{k=1}^n Y_k$. Then,

$$\sum \Pr(Y_k \neq X_k) = \sum \Pr(|X_1| > b_k) < \infty,$$

from which it follows that

$$(4) \quad b_n^{-1}(S_n - T_n) \xrightarrow{\text{a.s.}} 0$$

as $n \rightarrow \infty$. Also, for n sufficiently large,

$$E Y_n^2 \leq C b_n^2 \Pr(|X_1| > b_n),$$

and hence $\sum b_n^{-2} E Y_n^2 < \infty$. It then follows from a theorem of Kolmogorov (e.g. Loève [8, p. 238]) that

$$b_n^{-1}(T_n - ET_n) \xrightarrow{\text{a.s.}} 0$$

as $n \rightarrow \infty$ and consequently, from (4), that

$$b_n^{-1}(S_n - ET_n) \xrightarrow{\text{a.s.}} 0,$$

or equivalently (e.g. Loève [8, p. 247]),

$$b_n^{-1}(S_n - \text{med } S_n) \xrightarrow{\text{a.s.}} 0.$$

This again provides a contradiction and completes the proof of the theorem.

In view of the important condition (ii) of Theorem 1, it is interesting to investigate the oscillation behaviour of sums when the summands do not satisfy condition (ii). Unfortunately, it turns out that centering complications prevent us from giving a complete answer in the general case. These complications can, of course, be avoided by working with symmetric random variables and then it is possible to give a full explanation of the behaviour. This is what we shall do in the next theorem.

THEOREM 2. *Suppose the X_i are symmetric and that they do not belong to the domain of partial attraction of the normal distribution. If $\{b_n, n = 1, 2, 3, \dots\}$ is a monotone sequence of positive constants with $b_n \rightarrow \infty$ as $n \rightarrow \infty$, then for any $\epsilon > 0$, $\Pr(|S_n| > \epsilon b_n \text{ i.o.}) = 0$ or 1 according as $\sum \Pr(|X_1| > b_n)$ converges or diverges. Equivalently, $b_n^{-1} S_n \xrightarrow{\text{a.s.}} 0$ as $n \rightarrow \infty$ if and only if $\sum \Pr(|X_1| > b_n)$ converges.*

The proof of this theorem involves only minor modifications of the proof of Theorem 1 and will be omitted. The result is closely related to those of Feller [3].

REMARK. By way of comparison with the result of Theorem 2, it is worth noting that under the same conditions on the X_i , $b_n^{-1}S_n \xrightarrow{p} 0$ if and only if $n \Pr(|X_1| > \epsilon b_n) \rightarrow 0$ as $n \rightarrow \infty$ for every $\epsilon > 0$ ("p" denotes convergence in probability). This follows simply from the degenerate convergence criterion (e.g. Loève [8, p.317]) upon making use of the Lévy condition (2).

If further structural conditions are imposed on the X_i , it is often possible to investigate the matter of centering of the S_n process in sufficient detail to enable the symmetry assumption of Theorem 2 to be dispensed with while still obtaining the same behaviour. As an example of the possibilities, we have the following theorem.

THEOREM 3. *Suppose that the X_i belong to the domain of attraction of a nonnormal stable law of index $\alpha \neq 1$ which is not one-sided and that $EX_i = 0$ if $E|X_i| < \infty$. If $\{b_n, n = 1, 2, 3, \dots\}$ is a monotone sequence of positive constants with $b_n \rightarrow \infty$ as $n \rightarrow \infty$, then for any $\epsilon > 0$, $\Pr(|S_n| > \epsilon b_n \text{ i.o.}) = 0$ or 1 according as $\sum \Pr(|X_1| > b_n)$ converges or diverges. Equivalently, $b_n^{-1}S_n \xrightarrow{a.s.} 0$ as $n \rightarrow \infty$ if and only if $\sum \Pr(|X_1| > b_n)$ converges.*

PROOF. Again we can make use of a slightly modified version of the proof of Theorem 1 and in order to complete the proof of Theorem 3 it is just necessary to show that $b_n^{-1}ET_n \rightarrow 0$ as $n \rightarrow \infty$ where T_n is defined as in the proof of Theorem 1.

Firstly we note that under the conditions of the theorem we may write for $x > 0$, using Gnedenko and Kolmogorov [5, Theorem 2, p. 175],

$$\Pr(|X_1| > x) = M(x)x^{-\alpha},$$

where $M(x)$ is a function which is slowly varying as $x \rightarrow \infty$. It is necessary to treat separately the cases $\alpha < 1$ and $1 < \alpha < 2$.

If $\alpha < 1$,

$$\begin{aligned} |EY_n| &\leq E|Y_n| = \int_{|x| \leq b_n} |x| d \Pr(X_1 \leq x) \\ &\leq \int_0^{b_n} \Pr(|X_1| > x) dx, \end{aligned}$$

using integration by parts and, using standard properties of slowly varying functions (e.g. Feller [4, p. 273]), we have

$$\frac{\int_0^{b_n} \Pr(|X_1| > x) dx}{b_n \Pr(|X_1| > b_n)} \rightarrow \frac{1}{1 - \alpha}$$

as $n \rightarrow \infty$ so that we can choose a constant $c_1 > 0$ such that

$$(5) \quad |EY_n| \leq c_1 b_n \Pr(|X_1| > b_n)$$

for each n .

If, on the other hand, $1 < \alpha < 2$, we have since $EX_1 = 0$,

$$\begin{aligned} |EY_n| &= \left| \int_{|x| \leq b_n} x d \Pr(X_1 \leq x) \right| \\ &= \left| \int_{|x| > b_n} x d \Pr(X_1 \leq x) \right| \\ &\leq \int_{|x| > b_n} |x| d \Pr(X_1 \leq x) \\ &= - \int_{b_n}^{\infty} x d \Pr(|X_1| > x) \\ &= b_n \Pr(|X_1| > b_n) + \int_{b_n}^{\infty} \Pr(|X_1| > x) dx, \end{aligned}$$

and

$$\frac{\int_{b_n}^{\infty} \Pr(|X_1| > x) dx}{b_n \Pr(|X_1| > b_n)} \rightarrow \frac{1}{\alpha - 1}$$

as $n \rightarrow \infty$, so that we can choose a constant $c_2 > 0$ such that

$$(6) \quad |EY_n| \leq c_2 b_n \Pr(|X_1| > b_n).$$

Take $c = \max(c_1, c_2)$. Then, from (5) and (6) we have in general that

$$(7) \quad \sum_{n=1}^{\infty} b_n^{-1} |EY_n| \leq c \sum_{n=1}^{\infty} \Pr(|X_1| > b_n) < \infty,$$

and the result $b_n^{-1} EY_n \rightarrow 0$ follows from an application of the Kronecker Lemma to (7). This completes the sketch of the proof.

Theorems 2 and 3 provide a convenient tool for obtaining results of the type of Chover [2]. We immediately obtain from Theorem 3, for example, the following extension of the result of [2].

COROLLARY 1. *Suppose that the X_i belong to the domain of normal attraction of a nonnormal stable law of index $\alpha \neq 1$ which is not one-sided and that $EX_i = 0$ if $E|X_i| < \infty$. Then,*

$$\Pr(\limsup_{n \rightarrow \infty} |n^{-\alpha^{-1}} S_n|^{(\log \log n)^{-1}} = e^{\alpha^{-1}}) = 1.$$

(Note that in the case of normal attraction, there is a constant $c > 0$ such that $\Pr(|X_1| > x) \sim cx^{-\alpha}$ as $x \rightarrow \infty$.)

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AUSTRALIAN NATIONAL UNIVERSITY